

Lecture 4

Independence.

Fix a probability space (Ω, \mathcal{F}, P) .

Defⁿ: Let E_1, E_2, \dots be events (finite or countable).

We say that they are independent

if $P(E_{i_1} \cap E_{i_2} \cap \dots \cap E_{i_k}) = P(E_{i_1}) \cdot P(E_{i_2}) \cdot \dots \cdot P(E_{i_k})$

for all choices of $i_1 < i_2 < \dots < i_k$.

Example: Consider throw of a die:

$E_1 =$ number is $\leq 2 = \{1, 2\}$, $P(E_1) = \frac{1}{3}$

$E_2 =$ number is even = $\{2, 4, 6\}$, $P(E_2) = \frac{1}{2}$

Since $P(E_1 \cap E_2) = P(\{2\}) = \frac{1}{6} = \frac{1}{2} \cdot \frac{1}{3} = P(E_1)P(E_2)$

the events E_1, E_2 are independent.

Let $E_3 =$ number is $\leq 3 = \{1, 2, 3\}$.

Then $P(E_2 \cap E_3) = P(\{2, 3\}) = \frac{1}{6} \neq \frac{1}{2} \cdot \frac{1}{2} = P(E_2)P(E_3)$

So E_2 and E_3 are not independent.

Defⁿ Random variables X_1, X_2, \dots are said to be independent if for any choice of $i_1 < i_2 < \dots < i_k$ and Borel sets A_1, \dots, A_k , the events $\{X_{i_j} \in A_j\}$ ($j=1, 2, \dots, k$) are independent.

$$\text{That is, } P(\{X_{i_1} \in A_1\} \cap \{X_{i_2} \in A_2\} \cap \dots \cap \{X_{i_k} \in A_k\}) \\ = \prod_{j=1}^k P(\{X_{i_j} \in A_j\})$$

Defⁿ Sub σ -algebras $\mathcal{G}_1, \mathcal{G}_2, \dots$ are said to be independent if $P(G_{i_1} \cap G_{i_2} \cap \dots \cap G_{i_k}) = \prod_{j=1}^k P(G_{i_j})$ for all choices of indices $i_1 < i_2 < \dots < i_k$ and events $G_{i_j} \in \mathcal{G}_{i_j}$.

Remark: Independence of events and random variables are special cases of independence of σ -algebras.

1) Let E_1, E_2, \dots be events.

Set $\mathcal{G}_j = \{\emptyset, E_j, E_j^c, \Omega\} = \sigma(E_j)$. Then,

E_1, E_2, \dots independent $\Leftrightarrow \mathcal{G}_1, \mathcal{G}_2, \dots$ independent.

2) Let X_1, X_2, \dots be random variables.

Let $G_j = \sigma(X_j)$. Then,

X_1, X_2, \dots independent $\Leftrightarrow G_1, G_2, \dots$ independent.

Lemma Let G, \mathcal{H} be sub σ -algebras of \mathcal{F} .

and let \mathcal{I}, \mathcal{J} be π -systems that generate

G and \mathcal{H} : $\sigma(\mathcal{I}) = G$ and $\sigma(\mathcal{J}) = \mathcal{H}$.

Then G, \mathcal{H} are independent if and only if

$P(I \cap J) = P(I)P(J)$ for all $I \in \mathcal{I}, J \in \mathcal{J}$. (*)

Proof: " \Rightarrow " is clear as $\mathcal{I} \subseteq G, \mathcal{J} \subseteq \mathcal{H}$.

" \Leftarrow ": Assume (*) holds. Define

$P_I(H) = P(I \cap H)$ for all $H \in \mathcal{H}$.

This is a measure on (Ω, \mathcal{H}) :

$$\bullet P_I(\emptyset) = P(I \cap \emptyset) = P(\emptyset) = 0$$

$$\begin{aligned} \bullet P_I\left(\bigcup_{i=1}^{\infty} A_i\right) &= P\left(I \cap \bigcup_{i=1}^{\infty} A_i\right) \\ &= P\left(\bigcup_{i=1}^{\infty} I \cap A_i\right) = \sum_{i=1}^{\infty} P(I \cap A_i) \\ &= \sum_{i=1}^{\infty} P_I(A_i) \text{ for disjoint } A_i. \end{aligned}$$

Moreover, $P'_I(H) = P(I) \cdot P(H)$ is a measure on (Ω, \mathcal{H}) .

Since measures are uniquely determined by π -systems, $P'_I(H) = P_I(H)$ for all $H \in \mathcal{H}$ by using (*) over all $H \in \mathcal{J}$.

So $P(I \cap H) = P(I)P(H)$ for all $I \in \mathcal{I}$, $H \in \mathcal{H}$

Repeating the argument on the left gives

$P(G \cap H) = P(G)P(H)$ for all $G \in \mathcal{G}$, $H \in \mathcal{H}$. \square

Remark • $\{E_i\}$ is a π system.

• Events of the form $\{X_i \leq x\}$ form a π -system generating $\sigma(X_i)$. So, to verify that

X_1, X_2, \dots are independent it suffices to check

$P(X_{i_1} \leq x_1 \ \& \ X_{i_2} \leq x_2 \ \& \dots) = P(X_{i_1} \leq x_1)P(X_{i_2} \leq x_2) \dots$

for all i_1, i_2, \dots , and x_1, x_2, \dots

Second Borel-Cantelli lemma:

Assume that E_1, E_2, \dots are independent events and $\sum_{i=1}^{\infty} P(E_i) = \infty$. Then,

$$P(\limsup_{n \rightarrow \infty} E_n) = P(E_n \text{ occurs infinitely often}) = 1$$

Proof: Recall that $\limsup_{n \rightarrow \infty} E_n = \bigcap_{n \in \mathbb{N}} \bigcup_{m \geq n} E_m$.

Its complement is $\bigcup_{n \in \mathbb{N}} \bigcap_{m \geq n} E_m^c$.

$$\text{Now } P\left(\bigcap_{m \geq n} E_m^c\right) \leq P\left(\bigcap_{m=n}^M E_m^c\right) = \prod_{m=n}^M P(E_m^c)$$

$$= \prod_{m=n}^M (1 - P(E_m))$$

$$\leq \prod_{m=n}^M e^{-P(E_m)}$$

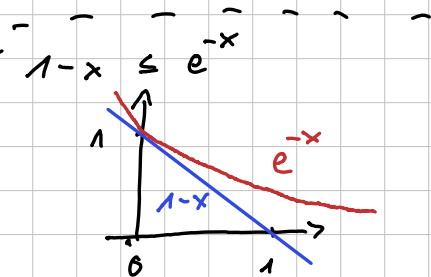
$$= \exp\left(-\sum_{m=n}^M P(E_m)\right)$$

$$\rightarrow \exp(-\infty) = 0.$$

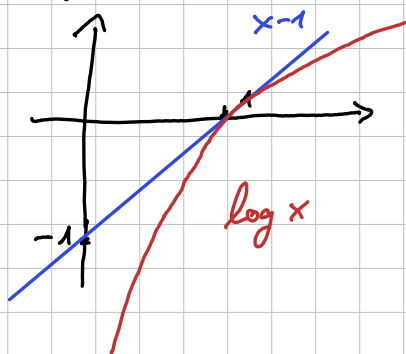
Hence $P\left(\bigcap_{m \geq n} E_m^c\right) = 0 \quad \forall n$

and by countable unions,

$$P\left(\bigcup_{n \in \mathbb{N}} \bigcap_{m \geq n} E_m^c\right) = 0. \quad \square$$



$$\log x \leq x-1$$



Remark: Independence is crucial!

If e.g. $E_1 = E_2 = \dots$, then $P(\limsup E_n) = P(E_n)$ which can take any value in $[0, 1]$.

Example: Take a random card from a deck of n cards on the n -th draw. Assume the cards are labelled $\{1, 2, \dots, n\}$. Let $E_n = \{\text{card 1 is drawn on } n\text{-th draw}\}$.

Assuming all draws are independent and uniform in probability, $P(E_n) = \frac{1}{n}$. Hence, by the second BC-lemma, $P(\text{"1 is drawn infinitely often"}) = P(\limsup_n E_n) = 1$, since $\sum P(E_n) = \sum \frac{1}{n} = \infty$.

Example: "Monkey & typewriter"

A monkey types a sequence of random characters on a keyboard. Assume each character has pos. prob. of occurring (at least $\epsilon > 0$). Let S be a fixed string of length s . Then,

$$\left\{ \begin{array}{l} P(\text{First } s \text{ characters are exactly } S) \geq \epsilon^s \\ P(\text{Characters } s+1, \dots, 2s \text{ are exactly } S) \geq \epsilon^s \\ \vdots \end{array} \right.$$

independent!

$$\sum_{k=1}^{\infty} P(k\text{-th set of } s \text{ characters is exactly } S) \geq \sum_{k=1}^{\infty} \epsilon^s = \infty.$$

By the second BC. lemma:

$$P(\text{monkey types } S \text{ infinitely often}) = 1.$$

Note: If E_1, E_2, \dots is a sequence of independent events, and since

$$\sum P(E_j) = \infty \quad \text{or} \quad \sum P(E_j) < \infty$$

must hold, the BC lemmas give

$$P\left(\limsup_n E_n\right) = \begin{cases} 0 & \sum P(E_n) < \infty \\ 1 & \sum P(E_n) = \infty \end{cases}.$$

This is a special case of Kolmogorov's 0-1 law!

Defⁿ Let X_1, X_2, \dots be a sequence of random variables. Set $\mathcal{T}_n = \sigma(X_{n+1}, X_{n+2}, \dots)$ and $\mathcal{T} = \bigcap_{n \in \mathbb{N}} \mathcal{T}_n$. We say \mathcal{T} is a tail σ -algebra.

Some "typical" events in the tail σ -algebra (tail events) are

$$\left\{ \lim X_n \text{ exists} \right\}, \left\{ \sum_{n=1}^{\infty} X_n \text{ converges} \right\}$$

Theorem (Kolmogorov 0-1 law)

Let X_1, X_2, \dots be independent random variables.

Then, for every $T \in \mathcal{T}$, either $P(T) = 0$ or 1 .

In particular, if \mathcal{E} is a \mathcal{T} measurable random variable, then $P(\mathcal{E} = c) = 1$ for some $c \in [-\infty, \infty]$.

Proof: 1) Define $\mathcal{X}_n = \sigma(X_1, \dots, X_n)$.

Note that \mathcal{X}_n and $\mathcal{T}_n = \sigma(X_{n+1}, \dots)$ are independent for all $n \in \mathbb{N}$.

2) Since $\mathcal{T} \subseteq \mathcal{T}_n$ for all n , \mathcal{T} and \mathcal{X}_n are independent.

3) $\mathcal{X}_\infty = \sigma(X_1, X_2, \dots)$ and \mathcal{T} are independent because $\bigcup_{n \geq 1} \mathcal{I}_n$ is a π system that generates \mathcal{X}_∞

4) $\mathcal{T} \subseteq \mathcal{X}_\infty$ and \mathcal{T} is independent of itself !!!

So for any $F \in \mathcal{T}$, $P(F \cap F) = P(F)P(F)$.

Hence $P(F) = x$ is a solution to $x = x^2$
 $\Rightarrow x \in \{0, 1\}$.

Now if ξ is \mathcal{T} measurable we set

$$c = \sup \{x : P(\xi \leq x) = 0\} \in [-\infty, \infty].$$

Then, $\{\xi < c\} = \bigcup_n \underbrace{\{\xi \leq c - \frac{1}{n}\}}_{\text{null set by def}} \Rightarrow P(\{\xi < c\}) = 0$

and

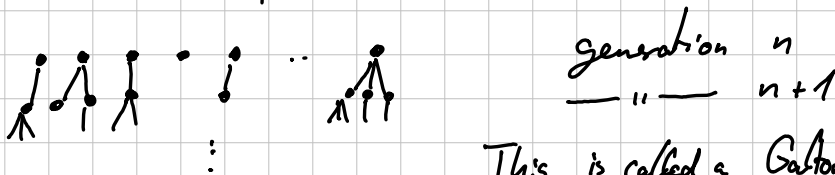
$$\{\xi \leq c\} = \bigcap_n \underbrace{\{\xi \leq c + \frac{1}{n}\}}_{\text{not null by def}} \Rightarrow P(\{\xi \leq c\}) = 1$$

so by 0-1 law almost sure

Hence $P(\{\xi = c\}) = 1$.

"Application" (Extra / non-examinable)

We consider a population (e.g. of humans) for which every member has offspring independent of other members & exactly one ancestor (e.g. matrilineal or patrilineal descent):



This is called a Galton-Watson process.

The following is a tail event: let v_n be an arbitrary member of the population at generation n . Then
 $\{ \text{a randomly picked individual in a population will eventually be descended from } v_n \} \in \mathcal{T}$.

Hence by the Kolmogorov 0-1 law, eventually everyone or no one will be descended from v_n .

NB: Recent studies into Y chromosomes & mitochondrial DNA suggest that the last time this happened was 280,000 years ago (patrilineal) and 160,000 years ago (matrilineal).